## STA414/2104

#### Statistical Methods for Machine Learning II

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#### Announcements

- HW 4 is released. Due on Apr 5, 2pm.
- TA OHs will be on the next week, time TBA.

# Last time

- Unsupervised learning
- Mixture models
- K-means
- EM Algorithm

#### Today

- PCA
- PPCA
- Autoencoders
- Recommender Systems

## Continuous Latent Variable Models

- Often there are some unknown underlying causes of the data.
- So far we have looked at models with discrete latent variables, such as mixture of Gaussians.
- Sometimes, it is more appropriate to think in terms of continuous factors which control the data we observe.
- Motivation: for many datasets, data points lie close to a manifold of much lower dimensionality compared to that of the original data space.
- Training continuous latent variable models often called dimensionality reduction, since there are typically many fewer latent dimensions.
- Examples: Principal Components Analysis, Factor Analysis, Independent Components Analysis

### Intrinsic Latent Dimensions

• What are the intrinsic latent dimensions in these two datasets?





• How can we find these latent dimensions from this high-dimensional data.

## Intrinsic Latent Dimensions

• In this dataset, there is only 3 degrees of freedom of variability, corresponding to vertical and horizontal translations, and the rotations.



- Each image undergoes a random displacement and rotation within some larger image field.
- The resulting images have 100 x 100 = 10,000 pixels.

### Generative View

- Each data example generated by first selecting a point from a distribution in the latent space, then generating a point from the conditional distribution in the input space
- Simplest latent variable models: Assume Gaussian distribution for both latent and observed variables.
- This leads to probabilistic formulation of the Principal Component Analysis.



- We will first look at standard PCA, and then consider its probabilistic formation.
- Advantages of probabilistic formulation: use of EM for parameter estimation, mixture of PCAs, Bayesian PCA.

### Latent dimensions

In practice, even though data is very high dimensional, important features can be accurately captured in a low dimensional subspace.



# Principal Component Analysis

- Used for data compression, visualization, feature extraction, dimensionality reduction.
- The goal is find M principal components underlying
   D-dimensional data
  - select the top M eigenvectors of S (data covariance matrix):  $\{\mathbf{u}_1, ..., \mathbf{u}_M\}$ .
  - project each input vector x into this subspace, e.g.





# Principal Component Analysis

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- Two views/derivations:
  - Maximize variance (scatter of green points).
  - Minimize error (red-green distance per data point).



#### Maximum Variance Formulation

• Consider a dataset  $\{x_1, ..., x_N\}$ , where  $x_n$  is in  $R^D$ . Our goal is to project data onto a

$$z_{n1} = \mathbf{x}_n^T \mathbf{u}_1.$$

 $\mathbf{u}_1^T \mathbf{u}_1 = 1.$ 



$$\frac{1}{N} \sum_{n=1}^{N} \frac{T}{1} n \qquad \begin{array}{cccc} T & 2 & T \\ 1 & n & 1 & 2 & 1 & 1 \\ & & & \\ \sum_{n=1}^{N} n & & & \\ & & & \\ & & & \sum_{n=1}^{N} |u_1^n| = 1 \quad n & T \\ \end{array}$$

$$\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1 + \lambda (1 - \mathbf{u}_1^T \mathbf{u}_1)$$

• Setting the derivative with respect to u<sub>1</sub> to zero:

$$\mathbf{S}\mathbf{u}_1 = \lambda_1 \mathbf{u}_1.$$

• Hence  $u_1$  must be an eigenvector of S.

$$\sum_{n=1}^{N} \sum_{n=1}^{T} \frac{1}{n} n = 1 \quad n = 1 \quad 1$$

$$\sum_{n=1}^{N} n = 1 \quad n = 1$$

$$\sum_{n=1}^{N} |u_1^n| = 1 \quad n = T$$

#### $\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1 + \lambda (1 - \mathbf{u}_1^T \mathbf{u}_1)$

• Setting the derivative with respect to u<sub>1</sub> to zero:

$$\mathbf{S}\mathbf{u}_1 = \lambda_1 \mathbf{u}_1.$$

- Hence u<sub>1</sub> must be an eigenvector of S.
- The maximum variance of the projected data is given by:

$$\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1 = \lambda_1.$$

• Optimal  $u_1$  is the first principal component (eigenvector with maximal eigenvalue).

• Introduce a complete orthonormal set of D-dimensional basis vectors:  $\{\mathbf{u}_1,...,\mathbf{u}_D\}$  :

$$\mathbf{u}_i^T \mathbf{u}_j = \delta_{ij}$$
. (1 if i=j, 0 otw)

• Without loss of generality, we can write:

$$\mathbf{x}_n = \sum_{i=1}^D \alpha_{ni} \mathbf{u}_i, \quad \alpha_{ni} = \mathbf{x}_n^T \mathbf{u}_i.$$

Rotation of the coordinate system to a new system defined by the basis vectors  $\boldsymbol{u}_{\text{i}}.$ 

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Rotation of the coordinate system to a new system defined by the basis vectors  $\mathbf{u}_{i}$ .

- Our goal is to represent data points by the projection into M-dimensional subspace (plus some distortion):
- Represent M-dim linear subspace by the first M of the basis vectors:

$$\tilde{\mathbf{x}}_n = \sum_{i=1}^M z_{ni} \mathbf{u}_i + \sum_{i=M+1}^D b_i \mathbf{u}_i.$$
Specific to a data point Shared across all data points

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• Represent M-dim linear subspace by the first M of the basis vectors:

$$\tilde{\mathbf{x}}_n = \sum_{i=1}^M z_{ni} \mathbf{u}_i + \sum_{i=M+1}^D b_i \mathbf{u}_i.$$

where  $z_{ni}$  depend on the particular data point and  $b_i$  are constants.

• Objective: minimize distortion with respect to u<sub>i</sub>, z<sub>ni</sub>, and b<sub>i</sub>.

$$J = \frac{1}{N} \sum_{n=1}^{N} ||\mathbf{x}_n - \tilde{\mathbf{x}}_n||^2.$$



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- Minimizing with respect to z<sub>nj</sub>, b<sub>j</sub>:
- Hence, the objective reduces to:

$$J = \frac{1}{N} \sum_{n=1}^{N} \sum_{i=M+1}^{D} (\mathbf{x}_n^T \mathbf{u}_i - \bar{\mathbf{x}}^T \mathbf{u}_i)^2 = \sum_{i=M+1}^{D} \mathbf{u}_i^T \mathbf{S} \mathbf{u}_i.$$

• Minimize distortion with respect to u<sub>i</sub>: constraint minimization problem:

$$J = \frac{1}{N} \sum_{n=1}^{N} ||\mathbf{x}_n - \tilde{\mathbf{x}}_n||^2 = \sum_{i=M+1}^{D} \mathbf{u}_i^T \mathbf{S} \mathbf{u}_i.$$

$$\mathbf{S}\mathbf{u}_i = \lambda_i \mathbf{u}_i.$$
  
En given by:  $J = \sum_{i=1}^{D} \lambda_i.$ 

• The distortion is then given by:

•If we use matrix notation for data:  $\mathbf{X} = [\mathbf{x}_1^T; ..; \mathbf{x}_N^T]$   $\tilde{\mathbf{X}} = [\tilde{\mathbf{x}}_1^T; ..; \tilde{\mathbf{x}}_N^T]$ 

i = M + 1

$$J = \frac{1}{N} \|\mathbf{X} - \tilde{\mathbf{X}}\|_F^2$$
 Frobenius norm

• PCA is finding the best rank-M approximation!

$$\mathbf{S} = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \bar{\mathbf{x}}) (\mathbf{x}_n - \bar{\mathbf{x}})^T$$



•We don't care about  $u_3$ . We want  $u_1$ .



#### PCA summary

•Observe D-dimensional N data points:  $\mathbf{X} \in \mathbb{R}^{N imes D}$ 

variance matrix:

$$(\mathbf{x}_n - \bar{\mathbf{x}})^T = \frac{1}{N} (\mathbf{X} - 1 \cdot \bar{\mathbf{x}}^T)^{\mathsf{T}} (\mathbf{X} - 1 \cdot \bar{\mathbf{x}}^T)$$

I by the top eigenvectors of the matrix S.  $\mathbf{U} \mathbf{\Lambda} \mathbf{U}^T$ s  $\mathbf{u}_1,...,\mathbf{u}_D$ 

$$j = 1, ..., M$$
  
 $j = M + 1, ..., D$ 

## Applications of PCA

• Run PCA on 2429 19x19 grayscale images (CBCL database)



- Data compression: We can get good reconstructions with only 3 components.
- Pre-processing: We can apply a standard classifier to latent representation -- PCA with 3 components obtains 79% accuracy on face/non-face discrimination in test data vs.
  76.8% for mixture of Gaussians with 84 components.
- Data visualization: by projecting the data onto the first two principal components.

#### Learned Basis

• Run PCA on 2429 19x19 grayscale images (CBCL database)



## PCA vs. Fisher's LDA

• A comparison of PCA with Fisher's LDA for linear dimensionality reduction.



- PCA chooses direction of maximum variance (magenta curve) leading to strong class overlap (unsupervised).
- LDA takes into account the class labels (supervised), leading to a projection into the green curve.

## PCA for High-Dimensional Data

- In some applications of PCA, the number of data points is smaller than the dimensionality of the data space, i.e. N<D.
- In so far, we need to find the eigenvectors of the D x D data covariance matrix S, which scales as  $O(D^3)$ .
- Direct application of PCA will often be computationally infeasible.
- Solution: Let X be the N x D **centered** data matrix. The corresponding eigenvector equation becomes:

$$\frac{1}{N}\mathbf{X}^T\mathbf{X}\mathbf{u}_i = \lambda_i \mathbf{u}_i.$$

• Pre-multiply by X:

$$\frac{1}{N}\mathbf{X}\mathbf{X}^T(\mathbf{X}\mathbf{u}_i) = \lambda_i(\mathbf{X}\mathbf{u}_i).$$

### PCA for High-Dimensional Data

• Define  $v_i = Xu_i$ , and hence we have:

$$\frac{1}{N}\mathbf{X}\mathbf{X}^T\mathbf{v}_i = \lambda_i \mathbf{v}_i.$$

- This is an eigenvector equation for the N x N matrix
- It has the same N-1 eigenvalues as the original data covariance matrix S (which itself has an additional D-N+1 zero eigenvalues).
- Computational cost scales as  $O(N^3)$  rather than  $O(D^3)$ .
- To determine eigenvectors, we multiply by  $X^T$ :

$$\left(\frac{1}{N}\mathbf{X}^T\mathbf{X}\right)(\mathbf{X}^T\mathbf{v}_i) = \lambda_i\mathbf{X}^T\mathbf{v}_i.$$

• Hence X<sup>T</sup> v<sub>i</sub> is an eigenvector of S with eigenvalue  $\,\lambda_i$  .

## Probabilistic PCA

- Probabilistic, generative view of data.
- Key properties of the probabilistic PCA (PPCA):
  - It represents a constrained form of the Gaussian distribution.
  - We can derive EM algorithm for PCA which is computationally efficient.
  - PPCA allows us to deal with missing values in the data set.
  - We can formulate mixture of PPCAs in a principled way.
  - PPCA forms the basis for a Bayesian PCA, in which the dimensionality of the principal subspace can be determined from the data.
  - The existence of a likelihood function allows direct comparisons with other probabilistic density models
  - PPCA can be used to model class conditional densities and hence it can be applied to classification problems.

### Proba

#### • Key assumptions:

- underlying latent M-dim variab Gaussian  $z_{nj}$ formulati
- linear re D-dim ob
- isotropic dimensio

- W

#### p

- $\bullet$  Hence the mean of x is a linear function of z governed by the D x M matrix W and the D-dim vector  $\mu$  .
- We will see that the columns of W span the principal subspace of the data space (Columns of W are the principal components,  $\sigma^2$  is sensor noise).



• Draw a value of the latent variable from its prior distribution:

$$\hat{z} \sim p(z)$$

• Draw a value for x from from an isotropic Gaussian distribution:

$$\hat{x} \sim p(\mathbf{x}|\hat{z}) = \mathcal{N}(\mathbf{x}|\mathbf{w}\hat{z} + \boldsymbol{\mu}, \sigma^2 I).$$



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#### Redundancy in Parameterization

• The marginal distribution is governed by parameters W,  $\mu$  ,  $\sigma^2$  :

$$p(\mathbf{x}) = \int_{\mathbf{z}} p(\mathbf{z}) p(\mathbf{x}|\mathbf{z}) d\mathbf{z} = \mathcal{N}(\mathbf{x}|\mu, \mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})$$

- Redundancy in parameterization: rotation of the latent space coordinates.
- Let R be an orthogonal matrix, then define a new matrix:

$$\tilde{\mathbf{W}} = \mathbf{W}\mathbf{R}, \qquad \mathbf{R}\mathbf{R}^T = \mathbf{I}.$$

• Then

$$\tilde{\mathbf{W}}\tilde{\mathbf{W}}^T = \mathbf{W}\mathbf{R}\mathbf{R}^T\mathbf{W}^T = \mathbf{W}\mathbf{W}^T.$$

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- There is a whole fam Work matrices and of which give rise to the same marginal distribution.
- Rotations within the latent space.

#### Joint De

• Joint density for PPCA, where x is D-d

• When evaluating marginal distribution, we need to invert a D x D matrix **C**, which can be expensive.

• Reduce O(D<sup>3</sup>) to O(M<sup>3</sup>) by applying matrix inversion lemma (Shenman Morrison formula – check wikipedia):

$$\mathbf{C}^{-1} = \sigma^{-1}\mathbf{I} - \sigma^{-2}\mathbf{W}(\mathbf{W}^T\mathbf{W} + \sigma^2\mathbf{I})^{-1}\mathbf{W}^T$$

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#### Posterior **Distribu**

• Inference in PPCA amounts to

 $p(\mathbf{z}|\mathbf{x}) = \mathcal{N}(\mathbf{z}|\mathbf{x})$  $\mathbf{m} = \mathbf{M}^{-1}\mathbf{W}^{T}(\mathbf{x} - \mathbf{v})$  $\mathbf{v} = \sigma^{2}\mathbf{M}^{-1}, -\mathbf{W}$ (Exercise) $\mathbf{M} = \mathbf{W}^{T}\mathbf{W} + \sigma^{2}\mathbf{l}.$ 

- Mean of inferred z is a linear map of centered x.
- Posterior variance does not depend on the input x at all.
- Remember:  $\mathbf{C} = \mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I}.$   $\mathbf{C}^{-1} = \sigma^{-1}\mathbf{I} - \sigma^{-2}\mathbf{W}(\mathbf{W}^T\mathbf{W} + \sigma^2\mathbf{I})^{-1}\mathbf{W}^T$





• W



- Hence PPCA is a constrained Gaussian model.
- We can fit model parameters using maximum likelihood.

#### Maximum

• Model parameters can be determined usin our latent variables):

$$L(\theta; \mathbf{X}) = \log p(\mathbf{X}|\theta) = \sum_{n} \log p(\mathbf{x}_{n}|\theta)$$
  
=  $-\frac{N}{2} \log |\mathbf{C}| - \frac{1}{2} \sum_{n} (\mathbf{x}_{n} - \mu) \mathbf{C}^{-1} (\mathbf{x}_{n} - \mu)^{T}$   
=  $-\frac{N}{2} \log |\mathbf{C}| - \frac{1}{2} Tr[\mathbf{C}^{-1} \sum_{n} (\mathbf{x}_{n} - \mu) (\mathbf{x}_{n} - \mu)^{T}] + \text{const}$   
-  $|| - \frac{-1}{2}$   
• Maximizing with respect to the mean?  $\mu_{ML} = \bar{\mathbf{x}}$ .

• We then have:

$$\log p(\mathbf{X}|\theta) = -\frac{N}{2} \log |\mathbf{C}| - \frac{1}{2} Tr[\mathbf{C}^{-1}\mathbf{S}] + \text{const.}$$

- Maximizing with respect to W and  $\sigma^2$  can be solved directly.

### Maximum Likelihood

• Objective:

$$\log p(\mathbf{X}|\theta) = -\frac{N}{2} \log |\mathbf{C}| - \frac{1}{2} Tr[\mathbf{C}^{-1}\mathbf{S}] + \text{const.}$$

- C is model covariance; S is sample data covariance.
- In other words, we are trying to make the constrained model covariance as close as possible to the observed covariance, where "close" means the trace of the ratio.

### Maximum Likelihood

• Objective:

$$\log p(\mathbf{X}|\theta) = -\frac{N}{2} \log |\mathbf{C}| - \frac{1}{2} Tr[\mathbf{C}^{-1}\mathbf{S}] + \text{const.}$$

• Maximizing with respect to W (derivation skipped in class):

$$\mathbf{W}_{ML} = \mathbf{U}_M (\mathbf{L}_M - \sigma^2 \mathbf{I})^{1/2} \mathbf{R},$$

where

- U<sub>M</sub> is a D x M matrix whose columns are given by the M principal eigenvectors of the data covariance matrix S.
- L<sub>M</sub> is the M x M diagonal matrix containing M largest eigenvalues of S.
- R is an arbitrary M x M orthogonal matrix.
- If the eigenvectors have been arranged in the order of decreasing values of the corresponding eigenvalues, then the columns of W define the principal subspace of standard PCA.

### Maximum Likelihood

• Objective:

$$\log p(\mathbf{X}|\theta) = -\frac{N}{2} \log |\mathbf{C}| - \frac{1}{2} Tr[\mathbf{C}^{-1}\mathbf{S}] + \text{const.}$$

• Maximizing with respect to :  $\sigma^2$ 

$$\sigma_{ML}^2 = \frac{1}{D-M} \sum_{i=M+1}^{D} \lambda_i,$$

which is the average variance associated with the discarded dimensions.

### EM

- Instead of solving directly, we can use dimensional datasets.
- The complete-data log-likelihood take

$$\log p(\mathbf{X}, \mathbf{Z} | \mu, \mathbf{W}, \sigma^2) = \sum_n [\log p(\mathbf{x}_n | \mathbf{z}_n) + \log p(\mathbf{z}_n)]$$

• E-step: compute expectation of complete log likelihood with respect to posterior of latent variables z, using current parameters.

- We need to derive  $\mathbb{E}[\mathbf{z}_n], \mathbb{E}[\mathbf{z}_n \mathbf{z}_n^T]$  with respect to the true posterior: p(z | X).
- M-step: maximize with respect to parameters W and  $\sigma^2$ .
- Appealing property: EM avoids direct O(ND<sup>2</sup>) construction of covariance matrix!
- Instead EM involves sums over data cases: O(NDM).



• We can der

to zero:

• MLE parameters are the same.

• Inferring the distribution over latent variables is easier: The posterior mean reduces to:

$$\lim_{\sigma^2 \to 0} (\mathbf{W}^T \mathbf{W} + \sigma \mathbf{I})^{-1} \mathbf{W}^T (\mathbf{x} - \boldsymbol{\mu}) = (\mathbf{W}^T \mathbf{W})^{-1} \mathbf{W}^T (\mathbf{x} - \boldsymbol{\mu}),$$

which represents an orthogonal projection of the data point onto the latent space – standard PCA.

• Posterior covariance goes to zero!

#### EM for PPCA

• EM algorithm for PCA.



#### Autoencoders

- Neural networks can also be used for nonlinear dimensionality reduction.
- This is achieved by having the same number of outputs as inputs. These models are called autoencoders.
- Consider a multilayer perceptron that has D inputs, D outputs, and M hidden units, with M<D.
- It is useful if we can squeeze the information through some kind of bottleneck.
- If we use a linear network (linear activation) this is very similar to Principal Components Analysis.



### Autoencoders and PCA

• Given an input x, its corresponding reconstruction is given by:

$$y_k(\mathbf{x}, \mathbf{w}) = \sum_{j=1}^M w_{kj}^{(2)} \sigma\left(\sum_{i=1}^D w_{ji}^{(1)} x_i\right), \quad k = 1, .., D.$$

• We can determine the network parameters w by minimizing the reconstruction error:

$$E(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} ||y(\mathbf{x}_n, \mathbf{w}) - \mathbf{x}_n||^2.$$

• If the hidden and output layers are linear, it will learn hidden units that are a linear function of the data and minimize the squared error.



• The M hidden units will span the same space as the first m principal components. The weight vectors may not be orthogonal.

### Deep Autoencoders

• We can put extra nonlinear hidden layers between the input and the bottleneck and between the bottleneck and the output.

- This gives nonlinear generalization of PCA.
- It should be very good for non-linear dimensionality reduction.
- The network can be trained by the minimization of the reconstruction error function.
- Much harder to train.



### Geometrical Interpretation

• Geometrical interpretation of the mappings performed by the network with 2 hidden layers for the case of D=3 and M=2 units in the middle layer.



• The mapping  $F_1$  defines a nonlinear projection of points in the original D-space into the M-dimensional subspace.

 $\bullet$  The mapping  $F_2$  maps from an M-dimensional space into D-dimensional space .

### Deep Autoencoders

- We can consider very deep autoencoders.
- There is an efficient way to learn these deep autoencoders



• By row: Real data, Deep autoencoder with a bottleneck of 30 linear units, and 30-d PCA.



### Deep Autoencoders

- We can consider very deep autoencoders.
- Similar model for MNIST handwritten digits:

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Real data 30-d deep autoencoder 30-d logistic PCA 30-d PCA

• Deep auto produces much better reconstructions.



### Class Structure of the Data

• Do the 30-D codes found by the deep autoencoder preserve the class structure of the data?

• Take the 30-D activity patterns in the code layer and display them in 2-D using a new form of non-linear multi-dimensional scaling (UNI-SNE).

• Will the learning find the natural classes?

### Class Structure of the Data

• Do the 30-D codes found by the deep autoencoder preserve the class structure of the data?





or term-frequency matrix:

$$\log(1 + M(doc, w)) \sim USV$$

$$U = |doc| \times d, S = d \times d, V = d \times |w|.$$
Autoencoder 2–D Topic Space
talk.politics.mideast
rec.sport.hockey



#### Reuters dataset

• Autoencoder: 2000-500-250-125-2



#### Recommender systems: Why?

- • YouTube 400 hours of video are uploaded to YouTube every minute
- amazon.ca



353 million products and 310 million users

83 million paying subscribers and streams about 35 million songs

Who cares about all these videos, products and songs? People may care only about a few  $\rightarrow$  Personalization: Connect users with content they may use/enjoy.

Recommender systems suggest items of interest and enjoyment to people based on their preferences

#### Some recommender systems in action



#### The Netflix problem

Movie recommendation: Users watch movies and rate them out of  $5 \bigstar$ .

User	Movie	Rating
•	Thor	* ☆ ☆ ☆ ☆
•	Chained	* * ☆ ☆ ☆
•	Frozen	$\bigstar\bigstar\bigstar\bigstar$
Ø	Chained	****
Ø	Bambi	****
Ü	Titanic	* * * ☆ ☆
Ö	Goodfellas	****
0	Dumbo	****
Ċ	Twilight	* * ☆ ☆ ☆
3	Frozen	****
<b>·</b>	Tangled	* ☆ ☆ ☆ ☆

Because users only rate a few items, one would like to infer their preference for unrated items

#### PCA as a matrix factorization

Ultimately, PCA with M principal components finds the *optimal* rank-M approximation of  $X \in \mathbb{R}^{N \times D}$ , in terms of error  $||X^{T} - UZ^{T}||_{F}^{2}$ .

#### $\min \|X^{\top} - UZ^{\top}\|_{F}^{2} \text{ over } Z \in \mathbb{R}^{N \times M}, U \in \mathbb{R}^{D \times M}.$

Note that the case M = D corresponds to the original matrix X.

• Can we do something similar for recommender systems?

#### PCA as matrix factorization of X



 $\bar{\mathbf{x}} = \mu + z_1 \mathbf{u_1} + z_2 \mathbf{u_2} + z_3 \mathbf{u_3} + \dots$ 

where the vectors  $\mathbf{u}_i$  are the principal components of the data matrix  $\mathbf{X}$  (the latent factors).

We can do the same for our ratings matrix R. Rating of movie

#### $\mathbf{\tilde{e}}$ = average user+ $z_1$ comedy user+ $z_2$ drama user+ $z_3$ action user+...

These latent factors are idealized, the real latent factors do not necessarily reveal these semantic concepts so clearly.

#### Matrix Completion

- We just saw that PCA gives the optimal low-rank matrix factorization.
- Two ways to generalize this:
  - ► 1) Consider when X is only partially observed.
    - ► A sparse 1000 × 1000 matrix with 50,000 observations (only 5% observed).
    - ► A rank 5 approximation requires only 10,000 parameters, so it's reasonable to fit this.
    - Unfortunately, no closed form solution.
  - ► 2) Impose structure on the factors. We can get lots of interesting models this way.

#### Matrix completion problem

Matrix completion problem: Transform the table into a N users by M movies matrix R



- Data: Users rate some movies. R<sub>user,movie</sub>. Very sparse
- Task: Finding missing data, e.g. for recommending new movies to users. Fill in the question marks
- Algorithms: Alternating Least Square method, Gradient Descent, Non-negative Matrix Factorization, low rank matrix Completion, etc.

#### Latent factor models

- In our current setting, latent factor models attempt to explain the ratings by characterizing both movies and users on a number of factors *K* inferred from the ratings patterns.
- That is, we seek representations for movies and users as vectors in ℝ<sup>K</sup> that can ultimately be translated to ratings.
- For simplicity, we can associate these factors (i.e. the dimensions of the vectors) with idealized concepts like
  - ► comedy
  - ► drama
  - ▶ action
  - But also uninterpretable dimensions

Can we use the sparse ratings matrix **R** to find these latent factors automatically?

#### Alternating least squares

- Let the representation of user *n* in the *K*-dimensional space be u<sub>n</sub> and the representation of movie *m* be z<sub>m</sub>
- Assume the rating user *n* gives to movie *m* is given by a dot product:  $R_{nm} \approx u_n^T z_m$
- In matrix form, if:

$$\mathbf{U} = \begin{bmatrix} - & \mathbf{u}_{1}^{\mathsf{T}} & - \\ \vdots & \vdots \\ - & \mathbf{u}_{N}^{\mathsf{T}} & - \end{bmatrix} \text{ and } \mathbf{Z}^{\mathsf{T}} = \begin{bmatrix} | & & | \\ \mathbf{z}_{1} & \dots & \mathbf{z}_{M} \\ | & & | \end{bmatrix}$$

then:  $R \approx UZ^{T}$ 

• This is a matrix factorization problem!

# Approach: Matrix factorization methods



#### Cost for Matrix Factorization for Recommender Systems

• Recall PCA: To enforce  $X^T \approx UZ^T$ , we minimized

$$\min_{\mathbf{U},\mathbf{Z}} \|\mathbf{X}^{\top} - \mathbf{U}\mathbf{Z}^{\top}\|_{\mathsf{F}}^{2} = \sum_{i,j} (x_{ji} - \mathbf{u}_{i}^{\top}\mathbf{z}_{j})^{2}$$

where  $\mathbf{u}_i$  and  $\mathbf{z}_i$  are the *i*-th rows of matrices U and Z, respectively.

• How do we enforce  $R \approx UZ^{T}$ 

► Try

$$\min_{\mathbf{U},\mathbf{Z}}\sum_{i,j}(R_{ij}-\mathbf{u}_i^{\mathsf{T}}\mathbf{z}_j)^2$$

• Most entries of **R** are missing!

#### Alternating least squares

- Let  $O = \{(n, m) : \text{ entry } (n, m) \text{ of matrix } R \text{ is observed}\}\$
- Using the squared error loss, a matrix factorization corresponds to solving

$$\min_{\mathbf{U},\mathbf{Z}} \frac{1}{2} \sum_{(n,m)\in O} \left( R_{nm} - \mathbf{u}_n^{\mathsf{T}} \mathbf{z}_m \right)^2$$

- The objective is non-convex in U and Z and in fact it's generally NP-hard to minimize the above cost function.
- As a function of either U or Z individually, the problem is convex and easy to optimize. We can use coordinate descent, just like with K-means and mixture models!

Alternating Least Squares (ALS): fix Z and optimize U, followed by fix U and optimize Z, and so on until convergence.

#### Alternating least squares

ALS for Matrix Completion algorithm

- 1. Initialize  $U \mbox{ and } Z \mbox{ randomly}$
- 2. repeat until convergence
- 3. for n = 1, .., N do

4. 
$$u_n = \left(\sum_{m:(n,m)\in O} z_m z_m^{\top}\right)^{-1} \sum_{m:(n,m)\in O} R_{nm} z_m$$

5. for 
$$m = 1, .., M$$
 do

6. 
$$\mathbf{z}_m = \left(\sum_{n:(n,m)\in O} \mathbf{u}_n \mathbf{u}_n^\top\right)^{-1} \sum_{n:(n,m)\in O} R_{nm} \mathbf{u}_n$$

#### Gradient descent method

- We can also do full gradient descent for matrix completion.
- Minimize *f*(U, Z) with GD. Both U, Z are variables. Gradient descent step:

$$\begin{bmatrix} \mathsf{U} \\ \mathsf{Z} \end{bmatrix} \leftarrow \begin{bmatrix} \mathsf{U} \\ \mathsf{Z} \end{bmatrix} - \alpha \nabla f(\mathsf{U},\mathsf{Z}) \tag{1}$$

• Computation of the gradient term per iteration is expensive if all the index pairs in the ratings matrix are considered and R is large (e.g. Netflix).

#### Stochastic gradient descent method

Stochastic gradient descent for matrix completion (recall SGD from lecture 5). Attempt to minimize  $f(U, Z) = \frac{1}{2} \sum_{(n,m) \in O} (R_{nm} - u_n^T z_m)^2$ . For a randomly chosen observed pair (n, m) in R, the SGD update:

$$\begin{bmatrix} \mathbf{u}_n \\ \mathbf{z}_m \end{bmatrix} \leftarrow \begin{bmatrix} \mathbf{u}_n \\ \mathbf{z}_m \end{bmatrix} - \alpha \begin{bmatrix} \left( R_{nm} - \mathbf{u}_n^{\mathsf{T}} \mathbf{z}_m \right) \mathbf{z}_m \\ \left( R_{nm} - \mathbf{u}_n^{\mathsf{T}} \mathbf{z}_m \right) \mathbf{u}_n \end{bmatrix}$$
(2)

Algorithm:

- 1. Initialize  ${\sf U}$  and  ${\sf Z}$
- 2. repeat until "convergence"
- 3. Randomly select a pair  $(n, m) \in O$  among observed elements of R

4. 
$$\mathbf{u}_n \leftarrow \mathbf{u}_n - \alpha \left( R_{nm} - \mathbf{u}_n^\top \mathbf{z}_m \right) \mathbf{z}_m$$

5. 
$$\mathbf{z}_m \leftarrow \mathbf{z}_m - \alpha \left( R_{nm} - \mathbf{u}_n^\top \mathbf{z}_m \right) \mathbf{u}_n$$

#### **Netflix Prize**

